

Supplementary Material for ‘Learning Tree Structure in Multi-Task Learning’

A. Basic Lemmas for Proving Theorem 3

Before presenting the proof for Theorem 3, we first prove some useful lemmas.

LEMMA 3. $\|\mathbf{C}\mathbf{W}_h^T\|_{1,2} \leq (m-1)\sqrt{d}\|\mathbf{W}_h\|_F$.

Proof: For any matrix $\mathbf{A} \in \mathbb{R}^{r_1 \times r_2}$, we have $\|\mathbf{A}\|_{1,2} \leq \sqrt{r}\|\mathbf{A}\|_F$, where $r \leq \min(r_1, r_2)$ denotes the rank of \mathbf{A} and the inequality holds due to the Cauchy-Schwarz inequality. Based on the definition of the matrix \mathbf{C} , we have

$$\begin{aligned} \|\mathbf{C}\mathbf{W}_h^T\|_{1,2} &= \frac{1}{2} \sum_{i=1}^m \sum_{j \neq i}^m \|\mathbf{w}_{h,i} - \mathbf{w}_{h,j}\|_2 \\ &\leq \frac{1}{2} \sum_{i=1}^m \sum_{j \neq i}^m (\|\mathbf{w}_{h,i}\|_2 + \|\mathbf{w}_{h,j}\|_2) \\ &= (m-1)\|\mathbf{W}_h^T\|_{1,2} \\ &\leq (m-1)\sqrt{d}\|\mathbf{W}_h\|_F, \end{aligned}$$

where the first inequality holds due to the triangular inequality for vector norms. So we complete the proof. \blacksquare

LEMMA 4. For any matrix pair $\mathbf{A}, \hat{\mathbf{A}} \in \mathbb{R}^{d \times m}$, we have

$$\|\mathbf{C}\mathbf{A}^T\|_{1,2} - \|\mathbf{C}\hat{\mathbf{A}}^T\|_{1,2} \leq \|(\mathbf{C}\mathbf{A}^T - \mathbf{C}\hat{\mathbf{A}}^T)^{E(\mathbf{A})}\|_{1,2}.$$

The proof of Lemma 4 is similar to that of Lemma 1 in [1] and hence we omit it here.

LEMMA 5. Assume that the training data is normalized to have zero mean and unit variance. For $h \in \mathbb{N}_H$, if the regularization parameter λ_h satisfies Eq. (19), then with probability at least $1 - \exp(-\frac{1}{2}(\delta - dm \log(1 + \frac{\delta}{dm})))$, for an optimal solution $\hat{\mathbf{W}} = \sum_{h=1}^H \hat{\mathbf{W}}_h$ of problem (3) and any $\mathbf{W} = \sum_{h=1}^H \mathbf{W}_h \in \mathbb{R}^{d \times m}$, where $\{\mathbf{W}_h\}_{h=1}^H$ satisfy the sequential constraints, we have

$$\begin{aligned} \frac{1}{mn} \|\mathbf{X}^T \text{vec}(\hat{\mathbf{W}}) - \text{vec}(\mathbf{F}^*)\|_2^2 &\leq \frac{1}{mn} \|\mathbf{X}^T \text{vec}(\mathbf{W}) - \text{vec}(\mathbf{F}^*)\|_2^2 \\ &+ (m-1)\sqrt{d} \sum_{h=1}^H \lambda_h (\theta_h + 1) \left\| (\hat{\mathbf{W}}_h - \mathbf{W}_h)^{D(\mathbf{W}_h)} \right\|_F. \end{aligned} \quad (25)$$

Proof. Since $\hat{\mathbf{W}}$ is an optimal solution of problem (3), $\{\hat{\mathbf{W}}_h\}_{h=1}^H$ satisfy the sequential constraints, and for any $\mathbf{W} = \sum_{h=1}^H \mathbf{W}_h$ satisfying the constraints too, we have

$$\begin{aligned} \frac{1}{mn} \sum_{i=1}^m \|\mathbf{X}_i^T \sum_{h=1}^H \hat{\mathbf{w}}_{h,i} - \mathbf{y}_i\|_2^2 \\ \leq \frac{1}{mn} \sum_{i=1}^m \|\mathbf{X}_i^T \sum_{h=1}^H \mathbf{w}_{h,i} - \mathbf{y}_i\|_2^2 + \sum_{h=1}^H \lambda_h \left(\|\mathbf{C}\mathbf{W}_h^T\|_{1,2} - \|\mathbf{C}\hat{\mathbf{W}}_h^T\|_{1,2} \right). \end{aligned}$$

By substituting $y_{ji} = (\mathbf{x}_j^{(i)})^T \mathbf{w}_i^* + \epsilon_{ji}$, $i \in \mathbb{N}_m$, $j \in \mathbb{N}_n$ into the above inequality, we can obtain

$$\begin{aligned} \frac{1}{mn} \sum_{i=1}^m \|\mathbf{X}_i^T \sum_{h=1}^H \hat{\mathbf{w}}_{h,i} - \mathbf{f}_i^*\|_2^2 &\leq \frac{1}{mn} \sum_{i=1}^m \|\mathbf{X}_i^T \sum_{h=1}^H \mathbf{w}_{h,i} - \mathbf{f}_i^*\|_2^2 \\ &+ \sum_{h=1}^H \lambda_h \left(\|\mathbf{C}\mathbf{W}_h^T\|_{1,2} - \|\mathbf{C}\hat{\mathbf{W}}_h^T\|_{1,2} \right) + \frac{2}{mn} \sum_{h=1}^H \langle \mathbf{Z}, \hat{\mathbf{W}}_h - \mathbf{W}_h \rangle, \end{aligned} \quad (26)$$

where $\mathbf{Z} = [\mathbf{X}_1 \epsilon_1, \dots, \mathbf{X}_m \epsilon_m] \in \mathbb{R}^{d \times m}$ with its (j, i) th entry computed as $z_{ji} = \sum_{k=1}^n x_{jk}^{(i)} \epsilon_{ki}$ and $x_{jk}^{(i)}$ denotes the (j, i) th entry in \mathbf{X}_i for the i th task. Since $\mathbf{x}_j^{(i)}$ is normalized to have zero mean and unit variance and $\epsilon_{ji} \sim \mathcal{N}(0, \sigma^2)$, we have

$$z_{ji} \sim \mathcal{N}(0, \sigma^2).$$

By defining a variable $v_{ji} = \frac{1}{\sigma} z_{ji}$, we can get that $v_{ji} \sim \mathcal{N}(0, 1)$. Thus we can get that a variable u with the definition as

$$u = \sum_{j=1}^d \sum_{i=1}^m v_{ji}^2 = \frac{1}{\sigma^2} \|\mathbf{Z}\|_F^2,$$

which follows a chi-squared distribution with the degree of freedom as md . According to the Wallace inequality [2], for any $\delta > 0$ we have

$$\Pr(u \geq md + \delta) \leq \exp\left(-\frac{1}{2} \left(\delta - md \log\left(1 + \frac{\delta}{md}\right) \right)\right).$$

Since $u = \frac{1}{\sigma^2} \|\mathbf{Z}\|_F^2$, we obtain that

$$\begin{aligned} \Pr\left(\frac{2}{mn} \|\mathbf{Z}\|_F \leq \frac{2\sigma}{mn} \sqrt{md + \delta}\right) \\ = \Pr(u \leq md + \delta) \\ \geq 1 - \exp\left(-\frac{1}{2} \left(\delta - md \log\left(1 + \frac{\delta}{md}\right) \right)\right). \end{aligned} \quad (27)$$

Based on Assumption 1 and Eq. (27), with probability at least $1 - \exp(-\frac{1}{2}(\delta - md \log(1 + \frac{\delta}{md})))$ we have

$$\begin{aligned} \frac{2}{mn} \sum_{h=1}^H \langle \mathbf{Z}, \hat{\mathbf{W}}_h - \mathbf{W}_h \rangle \\ \leq \frac{2}{mn} \|\mathbf{Z}\|_F \sum_{h=1}^H \|\hat{\mathbf{W}}_h - \mathbf{W}_h\|_F \\ \leq \frac{2\sigma}{mn} \sqrt{md + \delta} \sum_{h=1}^H \theta_h \left\| (\hat{\mathbf{W}}_h - \mathbf{W}_h)^{D(\mathbf{W}_h)} \right\|_F. \end{aligned} \quad (28)$$

Moreover, by using Lemma 3 and 4, we have

$$\begin{aligned} \|\mathbf{C}\mathbf{W}_h^T\|_{1,2} - \|\mathbf{C}\hat{\mathbf{W}}_h^T\|_{1,2} \\ \leq \left\| (\mathbf{C}\mathbf{W}_h^T - \mathbf{C}\hat{\mathbf{W}}_h^T)^{E(\mathbf{W}_h)} \right\|_{1,2} \\ \leq (m-1)\sqrt{d} \left\| (\mathbf{W}_h - \hat{\mathbf{W}}_h)^{D(\mathbf{W}_h)} \right\|_F. \end{aligned} \quad (29)$$

By combing Eqs. (26), (28), and (29), with probability at least $1 - \exp(-\frac{1}{2}(\delta - md \log(1 + \frac{\delta}{md})))$ we have

$$\begin{aligned} \frac{1}{mn} \|\mathbf{X}^T \text{vec}(\hat{\mathbf{W}}) - \text{vec}(\mathbf{F}^*)\|_2^2 &\leq \frac{1}{mn} \|\mathbf{X}^T \text{vec}(\mathbf{W}) - \text{vec}(\mathbf{F}^*)\|_2^2 \\ &+ \sum_{h=1}^H \left(\frac{2\sigma}{mn} \sqrt{md + \delta} \theta_h + (m-1)\sqrt{d} \lambda_h \right) \left\| (\hat{\mathbf{W}}_h - \mathbf{W}_h)^{D(\mathbf{W}_h)} \right\|_F. \end{aligned}$$

By plugging Eq. (19) into the above equation, we complete the proof. \blacksquare

B. Proof of Theorem 3

Proof. By making \mathbf{W}_h take value of \mathbf{W}_h^* for $h \in \mathbb{N}_H$ in Eq. (25), we obtain

$$\frac{1}{mn} \|\mathbf{X}^T \text{vec}(\hat{\mathbf{W}}) - \text{vec}(\mathbf{F}^*)\|_2^2 \leq (m-1)\sqrt{d} \sum_{h=1}^H \lambda_h (\theta_h + 1) \left\| \hat{\mathbf{W}}_h^{D(\mathbf{W}_h)} \right\|_F, \quad (30)$$

where $\Delta_h = \hat{\mathbf{W}}_h - \mathbf{W}_h^*$ and $\Delta = \sum_{h=1}^H \Delta_h$. Under Assumption 1, we have

$$\left\| \Delta_h^{D(\mathbf{W}_h)} \right\|_F \leq \frac{\left\| \mathbf{X}^T \text{vec}(\Delta) \right\|_2}{\beta_h \sqrt{mn}}. \quad (31)$$

By substituting Eq. (31) into Eq. (30), we obtain

$$\left\| \mathbf{X}^T \text{vec}(\Delta) \right\|_2 \leq (m-1) \sqrt{mnd\mathcal{C}}. \quad (32)$$

Therefore we can directly get Eq. (20) from Eq. (32). Since from Assumption 1, we have

$$\left\| \hat{\mathbf{W}}_h - \mathbf{W}_h^* \right\|_F = \theta_h \left\| \left(\hat{\mathbf{W}}_h - \mathbf{W}_h^* \right)^{D(\mathbf{W}_h)} \right\|_F,$$

$$\left\| \mathbf{C} \hat{\mathbf{W}}_h^T - \mathbf{C}(\mathbf{W}_h^*)^T \right\|_{1,2} = \gamma_h \left\| \left(\mathbf{C} \hat{\mathbf{W}}_h^T - \mathbf{C}(\mathbf{W}_h^*)^T \right)^{E(\mathbf{W}_h)} \right\|_{1,2}.$$

By combing Eqs. (29), (31), and (20), we can easily prove Eqs. (21) and (22).

To prove $\hat{E}_h = E(\mathbf{W}_h^*)$, we need to prove the following two statements:

$$\forall (i, j) \in \hat{E}_h \Rightarrow (i, j) \in E(\mathbf{W}_h^*), \quad (33)$$

$$\forall (i, j) \in E(\mathbf{W}_h^*) \Rightarrow (i, j) \in \hat{E}_h. \quad (34)$$

We first prove Eq. (33) by contradiction. Assume there exists a pair (i', j') such that $(i', j') \in \hat{E}_h$, but $(i', j') \notin E(\mathbf{W}_h^*)$. Then according to the definitions of \hat{E}_h and $E(\mathbf{W}_h^*)$, we have

$$\begin{aligned} \left\| \left(\mathbf{C} \hat{\mathbf{W}}_h^T - \mathbf{C}(\mathbf{W}_h^*)^T \right)^{(i', j')} \right\|_2 &= \left\| \left(\mathbf{C} \hat{\mathbf{W}}_h^T \right)^{(i', j')} \right\|_2 \\ &> \frac{\gamma_h (m-1)^2 d\mathcal{C}}{\beta_h}, \end{aligned}$$

which contradicts Eq. (22), hence we prove Eq. (33). Next we prove Eq. (34) by contradiction. Similarly, assume there exists $(i'', j'') \in E(\mathbf{W}_h^*)$, but $(i'', j'') \notin \hat{E}_h$. Since $(i'', j'') \notin \hat{E}_h$, based on the definition of \hat{E}_h in Eq. (24) we have

$$\left\| \left(\mathbf{C} \hat{\mathbf{W}}_h^T \right)^{(i'', j'')} \right\|_2 \leq \frac{\gamma_h (m-1)^2 d\mathcal{C}}{\beta_h}.$$

Furthermore, using the condition in Eq. (23), we have

$$\begin{aligned} &\left\| \left(\mathbf{C} \hat{\mathbf{W}}_h^T - \mathbf{C}(\mathbf{W}_h^*)^T \right)^{(i'', j'')} \right\|_2 \\ &\geq \left\| \left(\mathbf{C}(\mathbf{W}_h^*)^T \right)^{(i'', j'')} \right\|_2 - \left\| \left(\mathbf{C} \hat{\mathbf{W}}_h^T \right)^{(i'', j'')} \right\|_2 \\ &> \frac{\gamma_h (m-1)^2 d\mathcal{C}}{\beta_h}. \end{aligned}$$

which contradicts Eq. (22). So Eq. (34) is correct, which completes the proof. \blacksquare

References

- [1] P. Gong, J. Ye, and C. Zhang. Robust multi-task feature learning. In *Proceedings of the 18th ACM SIGKDD international conference on Knowledge discovery and data mining (KDD)*, pages 895–903, 2012.
- [2] D. L. Wallace. Bounds on normal approximations to student's and the chi-square distributions. *The Annals of Mathematical Statistics*, pages 1121–1130, 1959.